A Model Based Approach to Produce Residential or Commercial Property Price Indices

16th Ottawa Group **Meeting** Rio de Janeiro Brazil May, 2019





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HIGHLIGHTS

- **Hedonic Double Imputation Laspeyres House Price Indices**
- We Link Sold Properties
- **Create Pseuso Housing Units**
- Calculate Accurate Indices with Reduced Sample Sizes

Plot of our Longitudinal Data ■ Why some prices increase and other decrease?

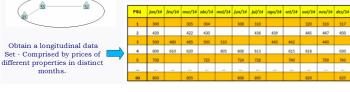
■ Which factors are related to changes in prices over time?

To answer this questions we apply

Mixed Effects Models

- Interesting technique to analyze longitudinal data because they offer us some prerrogatives:
- a) Analyze individual trajectories
- b) Identify variance components
- c) Predictors that explain intraindividual variance and variance among groups

Linking sold properties · We create PSEUDO HOUSING UNITS 0 1 2 3 4 5 6 J Time Table 1 - Prices of properties (US\$)



Methodology

- We define a fixed sample (S) 60 Specific Properties.
- We specify the Model Ex: To Calculate Results for Jan/2016
- We estimate Model coefficients taking into account data from the last 24 months
- Generate Predict Values for each property in the sample for Dec/15 and Jan/2016
- Calculate the Index for Jan/2016 $\hat{I}^{Jan16} = \frac{\sum \hat{P}_{icS}^{Jan16}}{\sum \hat{P}_{i.c}^{Dex15}}$

Model

■ Size, Month, Condo Characterístics, Neighbor (Zip Code), Distance to the sea;

 $Y_{ij} = \beta_0 + b_{0i} + \beta_1 T_{ij} + b_{1i} T_{ij} + \beta_2 B_i + \beta_3 P C_{ij} + \beta_4 D_{ij} + \epsilon_{ij}$ $b_{0i} \sim N(0, \sigma_0^2)$ $b_{1i} \sim N(0, \sigma_1^2)$

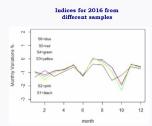
 $\epsilon_{ij} \sim N(0, \sigma_{\epsilon}^2)$

Model Estimation

| | Variable | Coef | Estimates | sd | p-value |
|--------------------------|----------------|-----------|-----------|---------|---------|
| neighborhood | Intercept | β_0 | 6378.54 | 685.37 | 0.000 |
| | Month | β_1 | -40.89 | 6.68 | 0.000 |
| | Area1 Barra | β_2 | - | - | - |
| | Area2 Bonsuc | β_2 | -2818.75 | 790.40 | 0.000 |
| | Area3 Freg | β_2 | -1459.82 | 988.11 | 0.145 |
| | Area4 Olaria | β_2 | -2945.64 | 1098.70 | 0.009 |
| | Area5 Penha | β_2 | -2788.51 | 903.31 | 0.003 |
| | Area6 Ramos | β_2 | -3579.34 | 840.71 | 0.000 |
| | Area7 V. Pen | β_2 | -2650.61 | 902.64 | 0.005 |
| condo characteristics | Pad Cond 1 | β_3 | - | - | - |
| | Pad Cond 2 | β_3 | 461.29 | 251.66 | 0.000 |
| | Pad Cond 3 | β_3 | 1160.72 | 270.71 | 0.000 |
| | Pad Cond 4 | β_3 | 1917.37 | 375.06 | 0.000 |
| | | 0 | | 667.27 | 0.000 |
| | Pad Cond 5 | β_3 | 4394.90 | 007.27 | 0.000 |
| | Dist Mar longe | β_4 | 4394.90 | - | - |
| Sea distance | | | | | |
| Sea distance | Dist Mar longe | β_4 | - | - | - |

Hedonic Double Imputation Laspeyres Price Indices 6 different Sample Sizes

| Mês | S1 | S2 | S3 | S4 | S5 | S6 |
|-------|----------|---------|---------|----------|---------|---------|
| Jan | -1.3616 | -1.2257 | -0.9809 | -0.9678 | -0.9665 | -0.9399 |
| Fev | -0.8593 | -0.8289 | -1.7305 | -1.5259 | -1.2960 | -1.1849 |
| Mar | -1.2923 | -1.2788 | -1.0337 | -0.9837 | -0.9106 | -0.8921 |
| Abr | -0.9461 | -0.9675 | -0.7374 | -0.8255 | -0.7968 | -0.7808 |
| Mai | -0.6097 | -0.3694 | -0.3694 | -0.4981 | -0.4772 | -0.4240 |
| Jun | -1.2654 | -1.2507 | -1.2477 | -1.3173 | -1.2772 | -1.2045 |
| Jul | -0.3982 | 0.1136 | 0.0307 | 0.0897 | 0.0275 | -0.0115 |
| Ago | -0.4451 | -0.3675 | -0.2513 | -0.2249 | -0.1297 | -0.0572 |
| Set | -1.5931 | -1.6205 | -1.1224 | -0.9200 | -0.6637 | -0.6117 |
| Out | -1.1872 | -1.2907 | -1.7590 | -2.3445 | -1.9898 | -1.8832 |
| Nov | -0.5819 | -0.6782 | -0.4999 | -0.3636 | -0.4039 | -0.3805 |
| Dez | -0.7385 | -0.6300 | -0.6215 | -0.6777 | -0.5903 | -0.5849 |
| Acum: | -10.7215 | -9.9261 | -9.8645 | -10.0846 | -9.0894 | -8.6112 |



Bootstrap Confidence Intervals for Sample S6

| Month | Variation | Lower | Upper |
|--------|-------------|------------|-------------|
| | (%) | | |
| Jan/16 | -0.93997576 | -1.0850158 | -0.79493576 |
| Fev/16 | -1.18498173 | -1.8455017 | -0.52446173 |
| Mar/16 | -0.89217714 | -1.0489771 | -0.73537714 |
| Abr/16 | -0.78083960 | -0.9905596 | -0.57111960 |
| Mai/16 | -0.42403444 | -0.6788344 | -0.16923444 |
| Jun/16 | -1.20455500 | -1.5397150 | -0.86939500 |
| Jul/16 | -0.01156223 | -0.3173222 | 0.29419777 |
| Ago/16 | -0.05720701 | -0.3806070 | 0.26619299 |
| Set/16 | -0.61173888 | -1.1605389 | -0.06293888 |
| Out/16 | -1.88321976 | -2.6201798 | -1.14625976 |
| Nov/16 | -0.38051861 | -0.6196386 | -0.14139861 |
| Dez/16 | -0.58491403 | -0.7652340 | -0.40459403 |

Concluding Remarks

- We calculate Quality Adjusted
- Longitudinal models allows more accurate results than other methods with the same sample
- Results based on transaction prices
- We are analysing alternative data sources and methods (Using Appraisal Prices)

Appendix

Mixed Effects Models

Extension of a standard linear model: $Y_{ij} = \beta_0 + \beta_1 X_{ij} + b_{0i} + b_{1i} X_{ij} + \epsilon_{ij}$ Matrix Notation: $Y_i = X_i \beta + Z_i b_i + \epsilon_i$ We can specify distributions: $b_i \sim NMV(0, \Sigma_i)$ $\varepsilon_i \sim NMV(0, R_i)$ $R_i = \sigma_i^2 h_n$ (cov Structure given by: $Z_i \Sigma_i Z_i' + R_i$. Intraindividual Variance: (R_i) Between Variance: $(Z_i \Sigma_i Z_i')$ PS: Covariance Pattern Models (CPMs): $Y_i = X_i \beta + \varepsilon_i$, where $c_i(X_i) = 0$.

 $cov(Y_i) = \Omega_i$

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